

## Appendix B

### The detail explanation of event reconstruction

#### B.1 Hit-PMT selection for vertex reconstruction

- (1) The hit channels in 200nsec timing window containing the maximum number of hits are selected. The reason we use 200nsec at first is that the maximum time of flight for a direct photon is approximately that.

- (2) The number of background hits in 200nsec window is estimated as follow;

$$N_{bg} = (t_3 - t_2) \times \frac{N_{hit}(t_1 : t_2) + N_{hit}(t_3 : t_4)}{(t_2 - t_1) + (t_4 - t_3)},$$

where  $N_{hit}(t_i : t_j)$  is number of hit-PMT in the timing between  $t_i$  and  $t_j$ , and  $t$  is as shown in Fig B.1,  $t_1$  and  $t_4$  are the beginning and end of the event, and  $t_2$  and  $t_3$  are the limits of the time window under consideration.

- (3) The maximum number of hits for each of 11 timing windows of width  $(200 \times n/11, n = 11, 10, \dots, 1)$  inside the 200nsec window determined above is calculated. For each window, the *significance* defined below is also calculated:

$$(\text{significance}) = \frac{N_{hit} - N_{bg}}{\sqrt{N_{bg}}}.$$

- (4) The window with maximum significance is selected.

- (5) It is checked whether any wider timing window than (4) with

$$(\text{significance}) > (\text{significance})_{\text{max}} \times 0.8$$

exist or not. If one does exist, all hit channels in this wider window are used for vertex reconstruction are used. If not the timing window which has maximum significance is taken. This is so that we can use as many signal hits as possible.

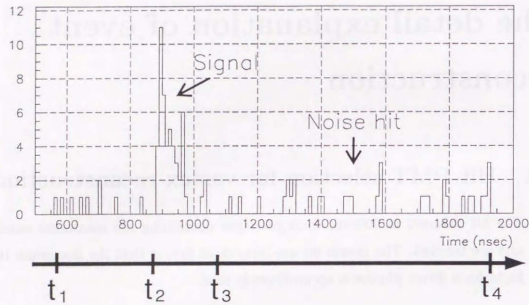


Figure B.1: The timing distribution of hit-PMTs in one event. The time, which is horizontal axis, is the relative timing of each hit-PMT obtained from ATM data.

## B.2 The definition of effective $N_{hit}$

The definition of effective  $N_{hit}$  ( $N_{eff}$ ) is as the following equation,

$$N_{eff} = \sum_{i=1}^{N_{hit}} \left[ (X_i + \epsilon_{tail} - \epsilon_{dark}) \times \frac{N_{all}}{N_{normal}} \times \frac{R_{cover}}{S(\theta_i, \phi_i)} \times \exp\left(\frac{r_i}{\lambda(\text{run})}\right) \times G_{time}(t) \right] \quad (\text{B.1})$$

The first factor is the correction to the each hit, where  $X_i$  is "occupancy" in order to estimate the effect of multi photo-electrons,  $\epsilon_{tail}$  is for the correction for late hits outside

the 50ns window, and  $\epsilon_{dark}$  is for dark noise correction. The definition of  $X_i$  is as follows:

$$X_i = \begin{cases} \frac{\lambda_i}{x_i} = \frac{\log \frac{1}{1-x_i}}{x_i} & x_i < 1 \\ 3 & x_i = 1 \end{cases} \quad (\text{B.2})$$

where  $x_i$  is the ratio of number of hit-PMTs to the number of PMTs around the  $i$ -th hit-PMT (normally nine ( $3 \times 3$ )) and  $\lambda_i$  is the estimated number of photons per one PMT in this  $3 \times 3$  PMTs region. This correction is to estimate the number of photons arrived at the  $i$ -th hit-PMT using the number of PMTs which does not hit in the  $3 \times 3$  PMTs region around the  $i$ -th hit-PMT. Here, the ratio of the PMTs which does not hit is measured to be  $1 - x_i$ . As while, when  $\lambda_i$  is the average number of photons per one PMT in the  $3 \times 3$  PMTs region, the probability not to hit is calculated by Poisson distribution as shown in the following equation,

$$P_0 = \frac{(\lambda_i)^0 \times e^{-\lambda_i}}{0!} = 1 - x_i. \quad (\text{B.3})$$

Fig B.2 shows the occupancy as a function of the number of hit-PMTs around one hit-PMT. When  $x_i = 1$ , the occupancy diverge, therefore, we determined that  $X_i = 3$  at  $x_i = 1$  by extrapolation as is seen in Fig B.2. It also minimizes the position dependence of  $N_{eff}$

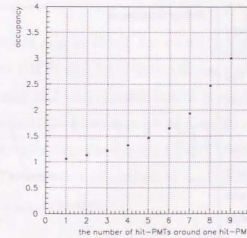


Figure B.2: The occupancy as a function of the number of hit-PMTs around one hit-PMT.

The second factor is the bad PMT correction, where  $N_{all}$  is the total number of PMTs, 11146, and  $N_{normal}$  is the number of properly operating PMTs for the relevant subrun.

The third factor is the effective photo coverage. The average of the photo coverage, which is the ratio of the area covered by PMT to all area, is 0.4041, ( $R_{cover}$ ). However, the effective photo coverage changes as the input angle of the photon to the PMT.



Therefore, we applied the coverage correction, where  $S(\theta_i, \phi_i)$  is the photo coverage from a view of  $\theta_i, \phi_i$  direction.

The fourth factor is the water transparency correction, where  $r_i$  is the distance from the reconstructed vertex to the  $i$ -th hit-PMT position,  $\lambda$  is water transparency as shown in Fig 4.12.

The fifth factor,  $G_{time}(i)$ , is the gain correction at single photo-electron level which depends on the time of the production of each PMT.

### B.3 Recent results of LINAC calibration

We recently took calibration data using an electron LINAC. This data taking was performed at 6 pipe positions and 7 beam energy points. The positions are  $(x, z) = (-4, 0), (-4, 12), (-8, 0), (-8, 12), (-12, 0)$  and  $(-12, 12)$  (m), all at  $y = -0.7$ m and the beam energies are shown in Table 5.2. Fig B.3 shows the difference of energy scale between data and Monte Carlo. The position dependence of the difference is less than 1% and the energy dependence is less than 0.3%. The differences of the energy, position and directional resolution between data and Monte Carlo are summarized in Table B.1.

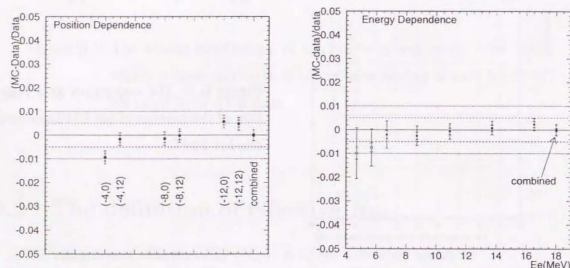


Figure B.3: The difference of energy scale between data and Monte Carlo.

The left figure shows the position dependence and the right one shows the energy dependence

Energy resolution	< 2%
Position resolution	< $\pm 5$ cm
Directional resolution	1.8 degree at 6.5MeV 0.3 degree at 16MeV

Table B.1: The differences of the energy resolution ((data-MC)/data), the position resolution (data-MC) and the directional resolution (data-MC) between data and Monte Carlo in new LINAC calibration data.

### B.4 Another algorithm of vertex reconstruction

As described in Section 7.2, we use a different algorithm of vertex reconstruction in order to reject the type of noise event shown in Fig 7.8. The essence of this reconstruction is selecting PMT hits which are considered to be caused by Cherenkov photons. For this purpose, we use the space-time correlation of hit-PMTs. The following criteria is applied in selecting PMTs: if a hit-PMT doesn't have neighboring hit-PMTs which satisfy the following condition,

$$\Delta r \leq r_{limit} \quad \text{and} \quad \Delta t \leq t_{limit},$$

we reject this hit-PMT for vertex reconstruction, where  $\Delta r$  is the spatial distance between PMTs and  $\Delta t$  is time difference between these PMTs.  $r_{limit}$  and  $t_{limit}$  are limits of distance and timing difference of given PMTs. These limits are chosen so that they reduce the dark noise as much as possible without reducing good PMT significantly. These limits are determined using LINAC data. Table B.2 shows the number of remained hit-PMTs in several conditions of  $r_{limit}$  and  $t_{limit}$ . For this analysis we take  $r_{limit} = 700$ cm and  $t_{limit} = 35$ ns.

Fig B.4 shows the distribution of the distance ( $\Delta l$ ) between the original vertex and the vertex from the new method. Some typical events with large  $\Delta l$  are shown in Fig 7.8. The vertex reconstruction of events with large  $\Delta l$  is generally clearly mistaken. Therefore, we use  $\Delta l$  for the event reduction as follows; events with  $\Delta l > 5$ m are rejected. The rejected events are classified by the number of small clusters as follows:  $\sim 24\%$  have two

$r_{limit}$ (cm)	$t_{limit}$ (nsec)	Number of hit-PMTs	Ratio (%)
400	20	40406	50
500	25	47907	60
600	30	53756	67
700	35	59054	74
800	40	63781	79

Table B.2: The number of remained hit-PMTs in several conditions to be neighboring PMTs, and  $r_{limit}$  and  $t_{limit}$  are given in the text. The number of all hit-PMTs is 80236. This is 5MeV LINAC data.

clusters; ~57% have one cluster; and ~19% have none.

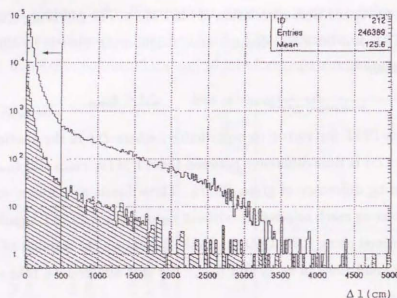


Figure B.4: The distribution of the distance between the original vertex and the vertex from the new method. These are the  $^8\text{B}$  Monte Carlo and the data after spallation cut (hatched region).

## B.5 The method of event clean up using vertex and direction

In addition to the criteria described above, we use the uniformity of the azimuthal angle distribution of hit-PMTs for identification of the event as shown in Fig 7.8.

The frame of reference has its center of the reconstructed vertex and the pole is the event direction. From this, the azimuth angle is calculated. Fig B.5 shows typical normal and noise events, and the azimuth angle distributions of each. The expected distribution is the dashed line, and we use the length of the arrow shown in the figure for the identification. If the event is normal, the hit-PMT distribution is uniform and the value is small, but if there is a cluster, the value is large.

Fig B.6 shows the scatter plot of this value and the difference between fitted and generated vertex in solar neutrino Monte Carlo events. This shows only mis-reconstructed events have large values.

Fig B.7 shows the distribution of the azimuthal angle non-uniformity in the final sample and in the solar neutrino Monte Carlo in two energy regions. The cut criteria is 0.4, and the efficiency is shown in the figure.

Fig B.8 shows the efficiency of the complete reduction combined with both methods described in Section B.4 and B.5 in real data and Monte Carlo. From the figure, we can see the real data and Monte Carlo are very consistent above our analysis threshold energy; the maximum of difference is 0.7%. This is taken as a systematic error in the solar neutrino analysis.





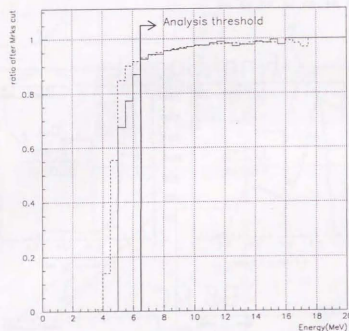


Figure B.8: The efficiency of the noise reduction as a function of energy. The solid line is real data of tightly chosen spallation events, and the dashed line is solar neutrino Monte Carlo.

## Appendix C

### Spallation events

#### C.1 Likelihood function of spallation event

As described in Section 7.2.4, three values ( $\Delta L$ ,  $Q_{res}$  and  $\Delta t$ ) are used for the spallation likelihood function. We calculate the probability density functions of  $\Delta L$ ,  $Q_{res}$  and  $\Delta t$  by spallation as  $F_{spa}(\Delta L)$ ,  $F_{spa}(Q_{res})$ ,  $F_{spa}(\Delta t)$ , and by events not caused by spallation as  $F_{non}(\Delta L)$ ,  $F_{non}(Q_{res})$ ,  $F_{non}(\Delta t)$ , respectively. Assuming that above three factors are independent, the spallation likelihood function ( $L_{spallation}$ ) is defined as follows,

$$L_{spallation} \equiv \frac{F_{spa}(\Delta L)}{F_{non}(\Delta L)} \cdot \frac{F_{spa}(Q_{res})}{F_{non}(Q_{res})} \cdot \frac{F_{spa}(\Delta t)}{F_{non}(\Delta t)} \quad (C.1)$$

For muons with tracks that cannot be well reconstructed,  $L_{spallation}$  is defined using only two parameters:

$$L_{spallation} \equiv \frac{F_{spa}(Q_{res})}{F_{non}(Q_{res})} \cdot \frac{F_{spa}(\Delta t)}{F_{non}(\Delta t)} \quad (C.2)$$

Spallation events have larger values of  $L_{spallation}$  than solar neutrino events.

Here, we describe how to find the probability density function. We gather samples of predominantly spallation events using the following selections,

- $\Delta t \leq 0.1 \text{ sec.}$  and  $N_{eff} \geq 50$  for  $\Delta L$
- $\Delta t \leq 0.1 \text{ sec.}$  and  $N_{eff} \geq 50$  for  $Q_{res}$
- $Q_{res} \leq 5.0 \times 10^5 \text{ p.e.}$  and  $\Delta L \leq 3 \text{ m}$  for  $\Delta t$

where we cut on  $N_{eff}$  because very low energy events are often caused by other background sources.

The distribution of  $\Delta L$  is made in six  $Q_{res}$  ranges as shown in Fig 7.11. In each histogram, the spallation component and phase space shown by dashed line ( $F_{non}(\Delta L)$ ) are included. Subtracting this phase space from the distribution, we get  $F_{spa}(\Delta L)$ .

The likelihood function for  $Q_{res}$  is defined the same way as for  $\Delta L$ . Fig C.1 shows the  $Q_{res}$  distribution of muons accompanied with a spallation event (a), and of muons without an associate spallation candidates (b). Subtracting (b) from (a), we can get the  $Q_{res}$  distribution from only the spallation component (c) ( $F_{spa}(Q_{res})$ ).

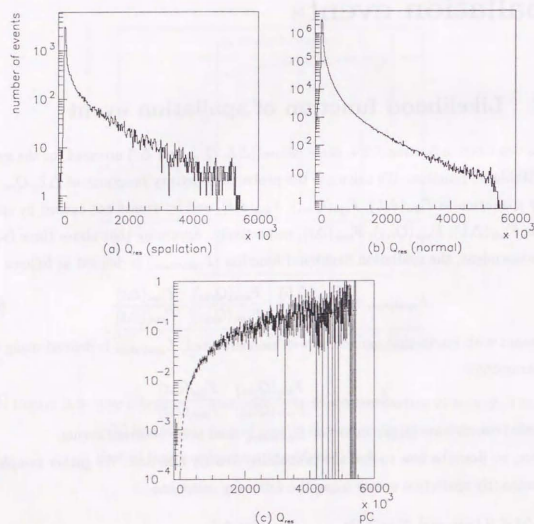


Figure C.1: The  $Q_{res}$  distribution of muons accompanied with a spallation event (a), of muons without any associated spallation candidates (b), and from only the spallation component (c), which is the differences (b) and (a).

Fig C.2 shows ( $\Delta t$ ) distribution. One can see a decay curve in correlation with the penetrating muons. This decay curve is the combination of many isotopes with different half-lives as shown in Table 7.1. For making the likelihood function for  $\Delta t$ , the decay curve as shown in Fig C.2 is fitted with the combination of isotopes with known half-life. The result is:

$$\frac{F_{spa}(\Delta t)}{F_{non}(\Delta t)} = 33900 \times 2^{-\frac{\Delta t}{0.011}} + 120100 \times 2^{-\frac{\Delta t}{0.0334}} + 338.6 \times 2^{-\frac{\Delta t}{0.0175}} + 1254 \times 2^{-\frac{\Delta t}{0.34}} \\ + 134.7 \times 2^{-\frac{\Delta t}{2.449}} + 676.1 \times 2^{-\frac{\Delta t}{7.134}} + 7.791 \times 2^{-\frac{\Delta t}{13.8}} \quad (C.3)$$

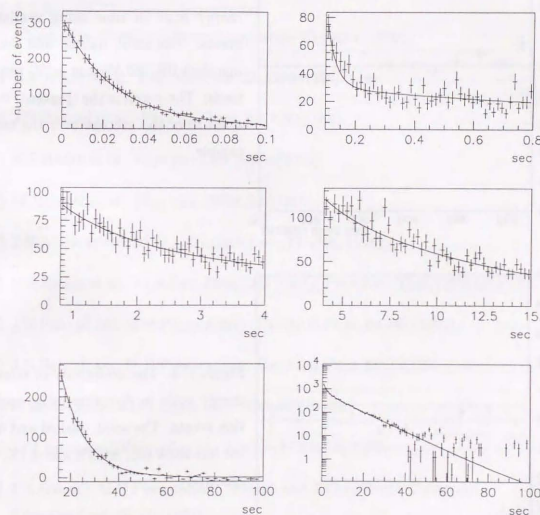


Figure C.2: The  $\Delta t$  distribution in several time range. The solid lines show fitting results.



## C.2 Spallation events as detector calibration

Spallation events are generated uniformly in volume, in direction and in time. Therefore, we can use these events as a good relative calibration source. Fig C.3 and Fig C.4 show the stability of relative energy scale in time and as a function of nadir, respectively. The uniformity of energy scale both in time and in direction are stable less than  $\pm 0.5\%$ .

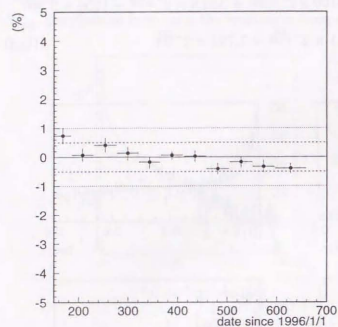


Figure C.3: The stability of relative energy scale in time using spallation events. The solid, dashed and dotted line show 0%,  $\pm 0.5\%$  and  $\pm 1\%$ , respectively. The y-axis is the (percent) deviation from the average over the entire sample.

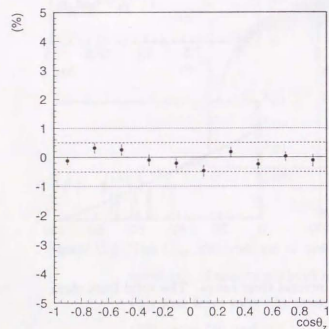


Figure C.4: The uniformity of relative energy scale in direction using spallation events. The solid, dashed and dotted line show 0%,  $\pm 0.5\%$  and  $\pm 1\%$ , respectively.

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